

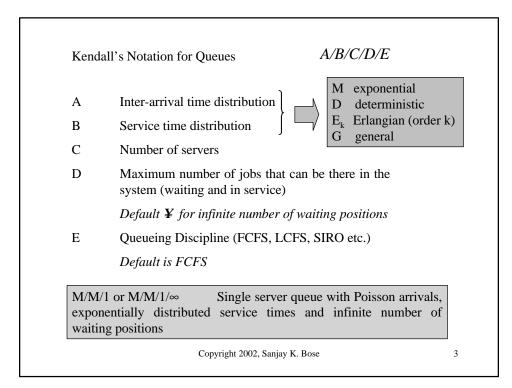
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Kendall's Notation for Queues

A/B/C/D/E

Shorthand notation where *A*, *B*, *C*, *D*, *E* describe the queue Applicable to a large number of simple queueing scenarios

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Little's Result

 $N = \mathbf{I}W \tag{2.9}$

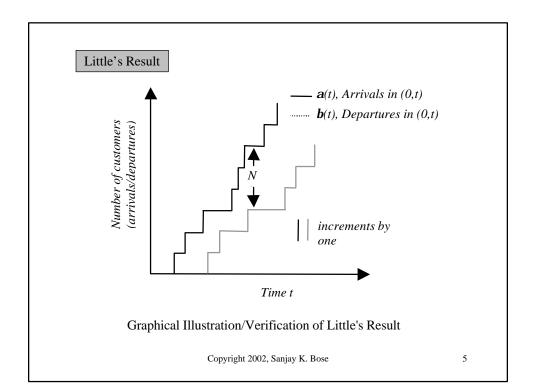
 $N_q = \mathbf{I}W_q \tag{2.10}$

Result holds in general for virtually all types of queueing situations where

 $I = Mean \ arrival \ rate \ of jobs \ that \ actually \ enter \ the \ system$

Jobs blocked and refused entry into the system will not be counted in I

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Little's Result

Consider the time interval (0,t) where t is large, i.e. $t \otimes \mathbf{Y}$

$$Area(t) = \text{area between } \boldsymbol{a}(t) \text{ and } \boldsymbol{b}(t) \text{ at time } t = \int_{0}^{t} [\boldsymbol{a}(t) - \boldsymbol{b}(t)] dt$$

Average Time W spent in system =
$$\lim_{t\to\infty} \frac{Area(t)}{a(t)}$$

Average Number
$$N$$
 in system = $\lim_{t\to\infty}\frac{Area(t)}{t}=\lim_{t\to\infty}\frac{\boldsymbol{a}(t)}{t}\frac{Area(t)}{\boldsymbol{a}(t)}$
Since, $\boldsymbol{l}=\lim_{t\to\infty}\frac{\boldsymbol{a}(t)}{t}$ Therefore, $N=\boldsymbol{l}W$

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The PASTA Property"Poisson Arrivals See Time Averages"

 $p_k(t) = P\{\text{system is in state } k \text{ at time } t \}$

 $q_k(t) = P\{\text{an arrival at time } t \text{ finds the system in state } k\}$

N(t) be the actual number in the system at time t

 $A(t, t+\mathbf{D}t)$ be the event of an arrival in the time interval $(t, t+\mathbf{D}t)$

$$\begin{split} q_k(t) &= \lim_{\Delta t \to 0} P \left\{ N(t) = k \mid A(t, t + \Delta t) \right\} \\ &= \lim_{\Delta t \to 0} \frac{P \left\{ A(t, t + \Delta t \mid N(t) = k \mid \right\} P \left\{ N(t) = k \right\}}{P \left\{ A(t, t + \Delta t) \right\}} = p_k(t) \end{split}$$

because $P\{A(t, t+Dt)/N(t) = k\} = P\{A(t, t+Dt)\}$

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7

Equilibrium Solutions for M/M/-/- Queues

Method 1: Obtain the differential-difference equations as in Section 1.2 or Section 2.2. Solve these under equilibrium conditions along with the normalization condition.

Method 2: Directly write the *flow balance equations* for proper choice of closed boundaries as illustrated in Section 2.2 and solve these along with the normalization condition.

Method 3: Identify the parameters of the birth-death Markov chain for the queue and directly use equations (2.7) and (2.8) as given in Section 2.2.

In the following, we have used this approach

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M/M/1 (or M/M/1/Y) Queue

For
$$\mathbf{r} < 1$$

$$\mathbf{m}_{k} = 0 \qquad k = 0$$

$$= \mathbf{m} \qquad k = 1, 2, 3, \dots$$

$$\mathbf{r} \qquad \mathbf{m}_{k} = \mathbf{r} \qquad \mathbf{$$

$$N = \sum_{i=0}^{\infty} i \boldsymbol{p}_i = \sum_{i=0}^{\infty} i \boldsymbol{r}^i (1 - \boldsymbol{r}) = \frac{\boldsymbol{r}}{1 - \boldsymbol{r}} \qquad W = \frac{N}{I} = \frac{1}{m(1 - \boldsymbol{r})} \qquad \text{Using Little's Result}$$

$$W_q = W - \frac{1}{m} = \frac{\mathbf{r}}{m(1-\mathbf{r})}$$
 $V_q = \mathbf{I}W_q = \frac{\mathbf{r}^2}{(1-\mathbf{r})}$ Using Little's Result

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9

M/M/1/∞ Queue with Discouraged Arrivals

$$\mathbf{I}_{k} = \frac{1}{k+1} \qquad \forall k
\mathbf{m}_{k} = 0 \qquad k = 0
= \mathbf{m} \qquad k = 1, 2, 3,$$
For $\mathbf{r} = \mathbf{I}/\mathbf{m} < \mathbf{Y}$

$$p_{k} = p_{0} \prod_{i=0}^{k-1} \frac{1}{\mathbf{m}(i+1)} = p_{0} \left(\frac{1}{\mathbf{m}}\right)^{k} \frac{1}{k!} \qquad (2.14)$$

$$p_{0} = \exp(-\frac{1}{\mathbf{m}}) \qquad (2.15)$$

$$N = \sum_{k=0}^{\infty} k p_{k} = \frac{1}{\mathbf{m}} \qquad \mathbf{I}_{eff} = \sum_{k=0}^{\infty} \mathbf{I}_{k} p_{k} = \mathbf{m} \left[1 - \exp(-\frac{1}{\mathbf{m}})\right]$$

$$N = \sum_{k=0}^{\infty} k p_k = \frac{1}{m}$$

$$V = \frac{N}{I_{eff}} = \frac{1}{m^2 \left[1 - \exp(-\frac{1}{m})\right]}$$

$$I_{eff} = \sum_{k=0}^{\infty} I_k p_k = m \left[1 - \exp(-\frac{1}{m})\right]$$

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M/M/1/∞ Queue with Discouraged Arrivals

In this case, *PASTA is not applicable* as the overall arrival process is not Poisson

$$\mathbf{p}_r = P\{\text{arriving customer sees } r \text{ in system} \qquad P\{E_i\} = p_i = e^{-1/m} \left(\frac{\mathbf{l}}{\mathbf{m}}\right)^i \frac{1}{i!}$$
(before joining the system)

DE be the event of an arrival in (t, t+Dt) E_i is the event of the system being in state i $P\{\Delta E \mid E_i\} = \frac{I\Delta t}{i+1}$

$$\mathbf{p}_{r} = P\{E_{r} \mid \Delta E\} = \frac{P\{E_{r}\}P\{\Delta E \mid E_{r}\}}{P\{\Delta E\}} = \frac{P\{E_{r}\}P\{\Delta E \mid E_{r}\}}{\sum_{i=0}^{\infty} P\{E_{i}\}P\{\Delta E \mid E_{i}\}}$$

$$\boldsymbol{p}_r = \left(\frac{1}{\mathbf{m}}\right)^{r+1} \frac{1}{(r+1)!} \left(\frac{e^{-1/\mathbf{m}}}{1 - e^{-1/\mathbf{m}}}\right) \quad W = \sum_{k=0}^{\infty} \frac{k+1}{\mathbf{m}} \boldsymbol{p}_k = \frac{1}{\mathbf{m}^2 (1 - e^{-1/\mathbf{m}})}$$

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11

M/M/m/∞ Queue (*m* servers, infinite number of waiting positions)

$$\mathbf{I}_k = \mathbf{I}$$
 $\forall k$ $\mathbf{m}_k = k\mathbf{m}$ $0 \le k \le (m-1)$
= $m\mathbf{m}$ $k \ge m$

For
$$\mathbf{r} = \mathbf{l}/\mathbf{m} < m$$

$$p_k = p_0 \frac{\mathbf{r}^k}{k!} \qquad \text{for } k \le m$$

$$= p_0 \frac{\mathbf{r}^k}{m! m^{k-m}} \qquad \text{for } k > m$$

$$(2.16)$$

Erlang's
$$p_0 = \left(\sum_{k=0}^{m-1} \frac{\mathbf{r}^k}{k!} + \frac{m\mathbf{r}^m}{m!(m-\mathbf{r})}\right)^{-1}$$
 (2.17)

$$P\{queueing\} = \sum_{k=m}^{\infty} p_k = C(m, \mathbf{r}) = p_0 \frac{m\mathbf{r}^m}{m!(m-\mathbf{r})}$$
(2.18)

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M/M/m/m Queue (m server loss system, no waiting)

$$egin{aligned} I_k &= I & k < m \\ &= 0 & otherwise & (Blocking or Loss Condition) \end{aligned}$$

$$\mathbf{m}_{k} = k\mathbf{m}$$
 $0 \le k \le m$
= 0 otherwise

For
$$\begin{cases} p_k = p_0 \frac{\mathbf{r}^k}{k!} & \text{for } k \le m \\ = 0 & \text{otherwise} \end{cases}$$
 (2.19)

For
$$\mathbf{r} = \frac{1}{\mathbf{m}} < \infty$$

$$p_k = p_0 \frac{\mathbf{r}^k}{k!} \qquad \text{for } k \le m$$

$$= 0 \qquad \text{otherwise}$$

$$p_0 = \frac{1}{\sum_{k=0}^m \frac{\mathbf{r}^k}{k!}} \qquad (2.19)$$

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13

M/M/m/m Queue (m server loss system, no waiting)

Simple model for a telephone exchange where a line is given only if one is available; otherwise the call is lost

Blocking Probability $B(m, \mathbf{r}) = P\{\text{an arrival finds all servers busy}\}$ and leaves without service}

$$B(m, \mathbf{r}) = p_0 \frac{\mathbf{r}^m}{m!}$$
 Erlang's B-Formula (2.21)

$$B(0, \mathbf{r}) = 1$$

$$B(m, \mathbf{r}) = \frac{\underline{rB(m-1, \mathbf{r})}}{1 + \frac{\underline{rB(m-1, \mathbf{r})}}{m}}$$
(2.22)

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M/M/1/K Queue (single server queue with K-1 waiting positions)

$$\mathbf{1}_k = \mathbf{1} \qquad \qquad k < K$$

$$\mathbf{m}_{k} = \mathbf{m}$$
 $k \le K$

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15

M/M/1/-/K Queue (single server, infinite number of waiting positions, finite customer population K)

$$\mathbf{1}_k = \mathbf{1}(K - k) \qquad k < K$$

$$\mathbf{m}_{k} = \mathbf{m}$$
 $k \le K$

For
$$p_k = p_0 \mathbf{r}^k \frac{K!}{(K-k)!}$$
 $k=1,...,K$ (2.25)

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Delay Analysis for a FCFS M/M/1/∞ Queue (Section 2.6.1)

Q: Queueing Delay (not counting service time for an arrival pdf $f_O(t)$, cdf $F_O(t)$, $L_O(s) = LT(f_O(t))$

W=Q+TW: Total Delay (waiting time and service time) for an arrival pdf $f_w(t)$, cdf $F_w(t)$, $L_w(s) = LT(f_w(t))$

> *T*: Service Time $f_T(t) = \mathbf{m}e^{-\mathbf{m}t}$ $F_T(t) = e^{-\mathbf{m}t}$ $L_T(s) = \frac{\mathbf{m}}{(s+\mathbf{m})}$

 $L_W(s) = \frac{\mathbf{m}}{(s+\mathbf{m})} L_Q(s) \quad f_W(t) = f_Q(t) * [\mathbf{m}e^{-\mathbf{m}}] \quad (2.30)$ Since $O \wedge T$

Knowing the distribution of either W or Q, the distribution of the other may be found

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17

For a particular arrival of interest -

$$F_O(t) = P\{\text{queueing delay} \le t\}$$

=
$$P\{\text{queueing time}=0\} + [\Sigma_{n^{3}I} P\{\text{queueing time} \leq t \mid \text{arrival found}]$$

$$n$$
 jobs in system $]p_n$

Erlang-n distribution for

$$F_{Q}(t) = (1 - \mathbf{r}) + (1 - \mathbf{r}) \sum_{n=1}^{\infty} \mathbf{r}^{n} \int_{x=0}^{t} \frac{\mathbf{m}(\mathbf{m}x)^{n-1}}{(n-1)!} e^{-\mathbf{m}x} dx$$
Example 1 distribution for sum of n exponential $\mathbf{r}.\mathbf{v}.\mathbf{s}$

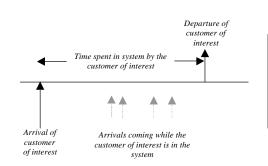
$$= (1 - \mathbf{r}) + (1 - \mathbf{r})\mathbf{r} \int_{0}^{t} \mathbf{m} e^{-\mathbf{m}x} \sum_{n=1}^{\infty} \frac{(\mathbf{m}x\mathbf{r})^{n-1}}{(n-1)!} dx$$
 (2.31)

$$= (1 - \mathbf{r}) + (1 - \mathbf{r})\mathbf{r} \int_{0}^{t} \mathbf{m} e^{-\mathbf{m}x(1-\mathbf{r})} dx = (1 - \mathbf{r}) + \mathbf{r}(1 - e^{-\mathbf{m}(1-\mathbf{r})})$$

$$f_{Q}(t) = \frac{dF_{Q}(t)}{dt} = \mathbf{d}(t)(1-\mathbf{r}) + \mathbf{I}(1-\mathbf{r})e^{-\mathbf{m}(1-\mathbf{r})}$$
(2.32)

$$f_W(t) = (1 - \mathbf{r}) m e^{-mt} + I(1 - \mathbf{r}) m \int_0^t e^{-m(1 - \mathbf{r})(t - x)} e^{-mx} dx = (m - I) e^{-(m - I)t}$$

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- PASTA applicable to this queue
- \bullet N and N_Q seen by an arrival same as the time-averaged values

Arrival/Departure of Customer/Job of Interest from a FCFS M/M/1Queue

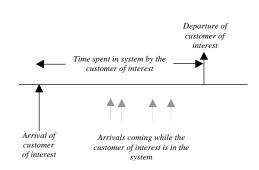
Let

 N^* = Number in the system that a job will see left behind when it departs

$$p_n*=P\{N^*=n\} \text{ for } N^*=0, 1,...., \mathbf{Y}$$

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19



For a FCFS queue, number left behind by a job will be equal to the number arriving while it is in the system.

$$G^{*}(z) = \sum_{n=0}^{\infty} z^{n} p_{n}^{*} = \sum_{n=0}^{\infty} z^{n} \int_{t=0}^{\infty} \frac{(\mathbf{I}t)^{n}}{n!} e^{-\mathbf{I}t} f_{W}(t) dt$$

$$= \int_{0}^{\infty} e^{-\mathbf{I}t(1-z)} f_{W}(t) dt = L_{W}(\mathbf{I} - \mathbf{I}z)$$

$$E\{N^{*}\} = \frac{dG^{*}(z)}{dz} \Big|_{z=1} = -\mathbf{I} \frac{dL_{W}(s)}{ds} \Big|_{s=0} = \mathbf{I}W$$
(2.36)

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An important general observation can also be made along the lines of Eq. (2.36).

Consider the number arriving from a Poisson process with rate I in a random time interval T where $L_T(s) = LT\{f_T(t)\}$. The generating function G(z) of this will be given by

$$G(z) = L_T(\mathbf{1} - \mathbf{1}z)$$

and the mean number will be $E\{N\}=1$ $E\{T\}$

This result will be found to be useful in various places in our subsequent analysis.

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21

Delay Analysis for the FCFS M/M/m/ Queue (Section 2.6.2)

Using an approach similar to that used for the M/M/1 queue, we obtain the following

$$f_{\mathcal{Q}}(t) = \left\{ 1 - p_0 \left[\frac{m \mathbf{r}^m}{m!(m - \mathbf{r})} \right] \right\} \mathbf{d}(t) + \left[\frac{m p_0 \mathbf{r}^m e^{-m(m - \mathbf{r})t}}{(m - 1)!} \right] u(t)$$
 (2.34)

$$f_{W}(t) = \left\{ 1 - p_{0} \left[\frac{m \mathbf{r}^{m}}{m!(m - \mathbf{r})} \right] \right\} \mathbf{m} e^{-\mathbf{m}} - \left[\frac{\mathbf{m} p_{0} \mathbf{r}^{m} [e^{-\mathbf{m}(m - \mathbf{r})t} - e^{-\mathbf{m}}]}{(m - 1)!(1 - m - \mathbf{r})} \right]$$
(2.35)

See Section 2.6.2 for the details and the intermediate steps

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