

**EE633 Queueing Systems (2015-16F)**  
**Solutions to the Mid-Term Examination**

**1. (a)**  $L_B(s) = \left(\frac{\mu}{s+\mu}\right)(0.5) + \left(\frac{\mu}{s+\mu}\right)(0.5)\left(\frac{\mu}{s+\mu}\right)(0.5) + \left(\frac{\mu}{s+\mu}\right)(0.5)\left(\frac{\mu}{s+\mu}\right)(0.5)\left(\frac{\mu}{s+\mu}\right)(0.5)L_B(s)$

$$L_B(s) \left[ 1 - \frac{0.25\mu^2}{(s+\mu)^2} \right] = \left(\frac{0.5\mu}{s+\mu}\right) \left[ 1 + \left(\frac{0.5\mu}{s+\mu}\right) \right]$$

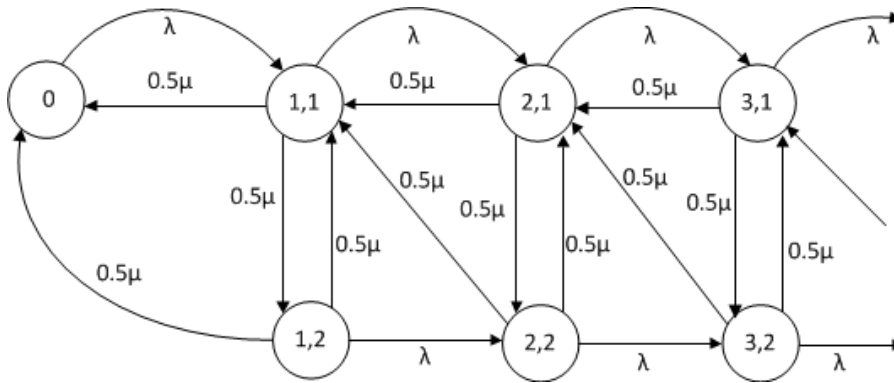
Therefore,  $L_B(s) = \frac{0.5\mu(s+1.5\mu)}{s^2 + 2s\mu + 0.75\mu^2} = \frac{0.5\mu}{s+0.5\mu}$

**(b)** As in (a), we can write  $\bar{X} = 0.5\left(\frac{1}{\mu}\right) + 0.25\left(\frac{2}{\mu}\right) + 0.25\left(\frac{2}{\mu} + \bar{X}\right)$

Therefore,  $\bar{X} = \frac{2}{\mu}$

**(c)** The queue will be stable if  $\lambda\bar{X} < 1$ , i.e.  $\frac{\lambda}{\mu} < 0.5$  or  $\rho < 0.5$

**(d) State Transition Diagram**



**(e)** Traffic offered is  $2\rho$ . Therefore, probability of server being idle will be  $(1-2\rho)$

**(f)** The Balance Equations can be written as follows –

$$p_{1,2}(\lambda + \mu) = p_{1,1}(0.5\mu)$$

$$p_{2,2}(\lambda + \mu) = p_{2,1}(0.5\mu) + p_{1,2}\lambda$$

$$p_{3,2}(\lambda + \mu) = p_{3,1}(0.5\mu) + p_{2,2}\lambda$$

$$p_{4,2}(\lambda + \mu) = p_{4,1}(0.5\mu) + p_{3,2}\lambda$$

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Equation Set 1

$$\lambda p_0 = 0.5\mu(p_{1,1} + p_{1,2})$$

$$\lambda(p_{1,1} + p_{1,2}) = 0.5\mu(p_{2,1} + p_{2,2})$$

$$\lambda(p_{2,1} + p_{2,2}) = 0.5\mu(p_{3,1} + p_{3,2})$$

$$\lambda(p_{3,1} + p_{3,2}) = 0.5\mu(p_{4,1} + p_{4,2})$$

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Equation Set 2

From Equation Set 1, we get –

$$(\lambda + \mu)P_2(z) = 0.5\mu P_1(z) + \lambda z P_2(z) \quad \Rightarrow \quad P_1(z) = 2(1 + \rho - \rho z)P_2(z) \quad \text{[A]}$$

From Equation Set 2, we get –

$$\lambda(p_0 + P_1(z) + P_2(z)) = \frac{\mu}{2z}(P_1(z) + P_2(z))$$

$$P_1(z) + P_2(z) = \frac{2\rho p_0 z}{1 - 2\rho z}$$

Note that  $p_0$  can be found using the Normalization Condition  $P_1(1) + P_2(1) + p_0 = 1$

Alternatively, we can directly use the result of part (e) to get  $p_0 = 1 - 2\rho$

$$\text{Therefore,} \quad P_1(z) + P_2(z) = (1 - 2\rho) \left( \frac{2\rho z}{1 - 2\rho z} \right) \quad \text{[B]}$$

Solving [A] and [B], we get –

$$P_2(z) = \frac{2\rho z(1 - 2\rho)}{(1 - 2\rho z)(3 + 2\rho - 2\rho z)}$$

$$P_1(z) = \frac{4\rho z(1 - 2\rho)(1 + \rho - \rho z)}{(1 - 2\rho z)(3 + 2\rho - 2\rho z)}$$

**[g]** Let  $\tilde{P}(z)$  be the generating function for the number in the system. This can be obtained

$$\text{directly from Equation [B] above as } \tilde{P}(z) = (1 - 2\rho) \left( \frac{2\rho z}{1 - 2\rho z} \right) + p_0 = \frac{1 - 2\rho}{1 - 2\rho z}$$

**[h]** Differentiating  $\tilde{P}(z)$  and evaluating at  $z=1$ , we get

$$N = \frac{2\rho}{1 - 2\rho}$$

**[i]** Applying Little's Result, we get  $W = \frac{2}{\mu(1 - 2\rho)}$

But we know from the results obtained for the M/G/1 queue that

$$W = \bar{X} + \frac{\lambda \overline{X^2}}{2(1 - \lambda \bar{X})}$$

Substituting appropriately from the results obtained earlier, we get

$$\frac{2}{\mu(1 - 2\rho)} = \frac{2}{\mu} + \frac{\lambda \overline{X^2}}{2(1 - 2\rho)}$$

$$\text{Therefore} \quad \overline{X^2} = \frac{8}{\mu^2}$$

[Differentiating  $L_B(s)$  twice and evaluating it at  $s=0$  gives the same result as above!]

**2. (a) Residual Life Approach**

$$P\{\text{first service is exceptional}\} = e^{-\lambda T}$$

$$\begin{aligned} \overline{BP} &= e^{-\lambda T} \left[ \overline{X^*} + (\lambda \overline{X^*}) \frac{\overline{X}}{1 - \lambda \overline{X}} \right] + (1 - e^{-\lambda T}) \left[ \frac{\overline{X}}{1 - \lambda \overline{X}} \right] \\ &= \frac{\overline{X}}{1 - \lambda \overline{X}} + e^{-\lambda T} \left[ \frac{\overline{X^*} - \overline{X}}{1 - \lambda \overline{X}} \right] = \frac{\overline{X} + \overline{\Delta} e^{-\lambda T}}{1 - \lambda \overline{X}} \end{aligned}$$

*Average Length of Busy Period*

$$\overline{T}_{cycle} = \frac{1}{\lambda} + \frac{\overline{X}}{1 - \lambda \overline{X}} + e^{-\lambda T} \left[ \frac{\overline{X^*} - \overline{X}}{1 - \lambda \overline{X}} \right] = \frac{1}{\lambda(1 - \lambda \overline{X})} + e^{-\lambda T} \left[ \frac{\overline{X^*} - \overline{X}}{1 - \lambda \overline{X}} \right]$$

*Average Cycle Length*

$$= \frac{1 + \lambda \overline{\Delta} e^{-\lambda T}}{\lambda(1 - \lambda \overline{X})} \quad \text{where } \overline{\Delta} = \overline{X^*} - \overline{X}$$

Note that if you consider a time interval of length  $t$ ,  $t \rightarrow \infty$ , then on the average it would have  $M = \lambda t$  arrivals and  $\frac{t}{T_{cycle}}$  Idle-Busy cycles of which  $N = e^{-\lambda T} \frac{t}{T_{cycle}}$  are ones with exceptional first service

**Mean Residual Time**

$$\begin{aligned} R &= \lim_{t \rightarrow \infty} \frac{1}{t} \left[ \sum_{i=1}^{M-N} \frac{1}{2} X_i^2 + \sum_{j=1}^N \frac{1}{2} X_j^{*2} \right] = \lim_{t \rightarrow \infty} \frac{1}{2} \left[ \frac{M-N}{t} \left( \frac{1}{M-N} \sum_{i=1}^{M-N} X_i^2 \right) + \frac{N}{t} \left( \frac{1}{N} \sum_{j=1}^N X_j^{*2} \right) \right] \\ &= \frac{1}{2} \left[ \lambda \overline{X^2} + \frac{\lambda e^{-\lambda T} (1 - \lambda \overline{X})}{1 + \lambda \overline{\Delta} e^{-\lambda T}} (\overline{X^{*2}} - \overline{X^2}) \right] \\ &= \frac{1}{2} \lambda \overline{X^2} \left[ 1 + \frac{e^{-\lambda T} (1 - \lambda \overline{X}) (\overline{X^{*2}} - \overline{X^2})}{1 + \lambda \overline{\Delta} e^{-\lambda T} \overline{X^2}} \right] \end{aligned}$$

**Mean Waiting Time in Queue**

$$W_q = \frac{R}{1 - \lambda \overline{X}} = \frac{1}{2} \frac{\lambda \overline{X^2}}{(1 - \lambda \overline{X})} \left[ 1 + \frac{e^{-\lambda T} (1 - \lambda \overline{X}) (\overline{X^{*2}} - \overline{X^2})}{1 + \lambda \overline{\Delta} e^{-\lambda T} \overline{X^2}} \right]$$

This was not asked but you can also easily find the probability  $p_0$  of the server being idle (or the system being empty) as  $p_0 = \frac{1/\lambda}{T_{cycle}} = \frac{1 - \lambda \overline{X}}{1 + \lambda \overline{\Delta} e^{-\lambda T}}$ . This would be useful to verify the results that you will get in part (b).

**(b) Imbedded Markov Chain Approach** (PASTA and Kleinrock's Principle are both applicable)

Following the usual notation, the state of the system at the departure instants will be -

$$\begin{aligned} n_{i+1} &= a_{i+1} & n_i &= 0, \text{ probability} = (1 - e^{-\lambda T}), \text{ idle period less than or equal to } T \\ &= a_{i+1}^* & n_i &= 0, \text{ probability} = e^{-\lambda T}, \text{ idle period more than } T \\ &= n_i + a_{i+1} - 1 & n_i &\geq 1 \end{aligned}$$

Since the system is considered at equilibrium, the subscripts can be dropped.

Taking expectations of both sides, we get

$$\bar{N} = p_0 \left[ (1 - e^{-\lambda T}) \lambda \bar{X} + e^{-\lambda T} \lambda \bar{X}^* \right] + \bar{N} - (1 - p_0) + (1 - p_0) \lambda \bar{X}$$

$$p_0 \left[ (1 - \lambda \bar{X}) + \left[ (1 - e^{-\lambda T}) \lambda \bar{X} + e^{-\lambda T} \lambda \bar{X}^* \right] \right] = 1 - \lambda \bar{X}$$

$$p_0 = \frac{1 - \lambda \bar{X}}{1 + \lambda \Delta e^{-\lambda T}} \quad \text{Same as before! Confirms result}$$

Raising both sides of the equation to the power of  $z$  and taking expectations, we get the following, where  $A(z) = L_B(\lambda - \lambda z)$  and  $A^*(z) = L_B^*(\lambda - \lambda z)$

$$\begin{aligned} P(z) &= A(z) \sum_{n=1}^{\infty} p_n z^{n-1} + p_0 \left[ e^{-\lambda T} A^*(z) + (1 - e^{-\lambda T}) A(z) \right] \\ &= \frac{A(z)}{z} [P(z) - p_0] + p_0 A(z) + p_0 e^{-\lambda T} [A^*(z) - A(z)] \end{aligned}$$

Therefore

$$P(z) = p_0 \frac{(1 - z)A(z) - z e^{-\lambda T} \{A^*(z) - A(z)\}}{A(z) - z}$$

**Note:** Answer is incomplete unless  $p_0$  is found somehow! Any method of finding it would be acceptable. Marks will be deducted if it is not found at all!